

## Lecture: Mar 3

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## 1 Hajek projection

The main idea is to use Hajek projections onto sets of the form:

$$\mathcal{S}_n = \left\{ \sum_{i=1}^n g_i(X_i) : g_i(X_i) \in L^2(P) \right\}$$

to approximate  $U_n$  by a sum of independent random variables.

### 1.1 Theorem

Let  $h$  be a symmetric kernel of order  $r$  with  $E[h^2(X_1, \dots, X_r)] < \infty$ . Let  $U_n$  be the associated U-statistic and  $\theta = E[U_n]$ . If  $\hat{U}_n$  is the projection of  $U_n - \theta$  onto  $\mathcal{S}_n$ , then

$$\hat{U}_n = \sum_{i=1}^n E[U_n - \theta | X_i] = \frac{r}{n} \sum_{i=1}^n \hat{h}_1(X_i),$$

where  $\hat{h}_1(x) = E[h(x, X_2, \dots, X_r)] - \theta$ .

**Proof:** The first equality is just a direct application of the last result from the previous lecture, noting that  $E[U_n - \theta] = 0$ . To show the second equality, we note that

$$E[h(X_\beta) - \theta | X_i] = \begin{cases} 0, & \text{if } i \notin \beta, \\ \hat{h}_1(X_i), & \text{if } i \in \beta. \end{cases}$$

Thus we get

$$\begin{aligned} E[U_n - \theta | x_i] &= \binom{n}{r}^{-1} \sum_{\beta \subset [n], |\beta|=r, i \in \beta} \hat{h}_1(X_i) \\ &= \binom{n}{r}^{-1} \binom{n-1}{r-1} h_1(X_i) = \frac{r}{n} \hat{h}_1(X_i), \end{aligned}$$

which implies the result.

### 1.2 Theorem

We have

1.  $\sqrt{n}\hat{U}_n \rightarrow^d N(0, r^2\xi_1)$ .
2.  $\sqrt{n}(U_n - \theta - \hat{U}_n) \rightarrow^p 0$ .
3.  $\sqrt{n}(U_n - \theta) = \sqrt{n}\hat{U}_n + o_p(1) \rightarrow^d N(0, r^2\xi_1)$ .

**Proof:** The first result follows from CLT. As

$$\text{var}(U_n) = \frac{r^2}{n} \xi_1 + o(n^{-2})$$

and

$$\text{var}(\hat{U}_n) = \frac{r^2}{n} \xi_1,$$

we have

$$\frac{\text{var}(\sqrt{n}(U_n - \theta))}{\text{var}(\sqrt{n}\hat{U}_n)} \rightarrow 1.$$

Using a result before, we have

$$\frac{\sqrt{n}(U_n - \theta)}{r\xi_1} - \frac{\sqrt{n}\hat{U}_n}{r\xi_1} \xrightarrow{p} 0.$$

By application of Slutsky's theorem, we can conclude the desired results.